Warmer and wetter or warmer and dryer? Observed versus simulated covariability of Holocene temperature and rainfall in Asia

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Abstract

Temperatures in Asia, and globally, are very likely to increase with greenhouse gas emissions, but future projections of rainfall are far more uncertain. Here we investigate the linkage between temperature and precipitation in Asia on interannual to multicentennial timescales using instrumental data, late Holocene paleoclimate proxy data and climate model simulations. We find that in the instrumental and proxy data, the relationship between temperature and precipitation is timescale-dependent. While on annual to decadal timescales, negative correlations dominate and thus cool summers tend to be rainy summers, on longer timescales precipitation and temperature are positively correlated; cool centuries tend to be dryer centuries in monsoonal Asia. In contrast, the analyzed CMIP5/PMIP3 climate model simulations show a negative correlation between precipitation and temperature on all timescales. Although many uncertainties exist in the interpre-

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tation of the proxy data, there is consistency between them and the instrumental evidence. This, and the persistence of the result across independent proxy datasets, suggests that the climate model simulations might be considerably biased, overestimating the short-term negative associations between regional rainfall and temperature and lacking long-term positive relationships between them.

Keywords: Asian summer monsoon, temperature, precipitation, climate variability

1 1. Introduction

The Asian summer monsoon winds transfer moisture from the tropical oceans onshore and release it as they cool while traveling inland, driven mainly by the thermal gradient between the surrounding oceans and the land surface (Fig. 1 and Turner and Annamalai [44]). The state and fate of the monsoon is of particular importance to the agricultural economies across Asia, yet, globally and across Asia, precipitation projections are far more uncertain than those for temperature [15].

Simulations of future precipitation in the Coupled Model Intercompari-9 son Project Phase 3 [CMIP3, 27] showed no consistent response in Asia to 10 increasing temperatures. The models in the more recent CMIP5 ensemble 11 [40] largely agree on an increase in rainfall amount and variability [28, 36]. 12 Nevertheless, the skill of the models in representing key features of the Asian 13 summer monsoon, such as its onset timing, duration and intensity has not 14 improved significantly from CMIP3 to CMIP5 [28, 36]. Improved consis-15 tency across models therefore does not guarantee improved future prediction 16



Figure 1: Overview of the study area and the dominant summer (orange) and winter (light blue and yellow) wind systems. Symbols show the paleoclimate proxy data sites and meteorological stations for which Table 1 and Supplementary Table 2 give more details.

¹⁷ skill, as many models have difficulties in simulating monsoon rainfall and
¹⁸ variability [22, 44].

In theory, global rainfall is likely to increase in a warmer world, as the 19 partial pressure of water vapor at saturation increases by $\sim 7\%$ per 1°C tem-20 perature increase, following the Clausius-Clapevron relationship [15]. Lo-21 cally, precipitation responses are difficult to project, as it is unclear if the 22 atmospheric pathways which relay evaporated oceanic moisture onto the con-23 tinents remain the same in a warmer atmosphere with greenhouse gas, aerosol 24 load, regional vegetation and land use changes. Analyzing trends of the last 25 50 years showed a warming but no consistent precipitation change across 26 Asia [44]. The thermal response to greenhouse gas forcing is better known 27 than the hydrological response. Thus, complementary information on future 28 rainfall can be gained by analyzing the relationship of precipitation and tem-29 perature (T-P relationship) on observational datasets. Direct extrapolation 30 of results based on largely naturally forced past temperature variability onto 31 a future where temperature changes are dominated by anthropogenic forc-32 ing, however, needs to be treated with caution, as the monsoon circulation 33 response may also be specific to the forcing, rather than temperature changes. 34 35

On daily to interannual timescales, negative correlations between local temperatures and precipitation in Asia were estimated from satellite and station data as well as from model simulations [1, 3, 42, 46]. This evident negative correlation between local temperature and precipitation roots in fundamental aspects of the hydrological cycle: rainy days tend to have a higher cloud cover and soil moisture, and thus lower temperatures through

insolation shielding and evaporative cooling, hot days are more likely to be 42 dry [3, 46, and references therein]. Over land, this anticorrelation was found 43 to be strongest in the summer months but persisted throughout the year. 44 On daily timescales, Williams et al. [46] observed differences to the monthly 45 analysis of Trenberth [42] and concluded that a timescale dependency of pro-46 cesses influencing the T-P relationship is already relevant between daily and 47 monthly scales. Due to the shortness of the observational record, however, 48 station- or satellite-based correlation studies are mostly limited to shorter 49 than decadal timescales. 50

Within monsoonal Asia, slow processes acting on interannual to centen-51 nial scales are likely to modify the boundary conditions for the monsoon 52 circulation, modulating its intensity, duration and distribution. Most of 53 them result in a positive association between regional temperatures, and 54 rainfall amounts: On the oceanic side, interannual to centennial precipi-55 tation changes in monsoonal Asia have been attributed to warmer surface 56 temperatures in the subtropical Pacific and the Indian Ocean, the source 57 areas of monsoonal moisture [42]. In the atmosphere, reducing (increasing) 58 the albedo of the Tibetan Plateau by lower (higher) snow cover in a warming 59 scenario, was proposed to increase (decrease) monsoonal intensity by damp-60 ing (strengthening) its role as an amplifying elevated heat source [50]. On 61 centennial timescales, proxy data suggests a wetter summer monsoon during 62 the warm Medieval Climate Anomaly, and a weakening during the cold pe-63 riod thereafter [e.g. 7, 33, 49]. This is consistent with the notion that the 64 Intertropical Convergence zone extends further north in warm periods than 65 in cold periods [34]. 66

On timescales of decades to centuries the nature and timescale-dependency of the T-P relationship within Asia and beyond is far from being understood. Here, we provide a systematic investigation of the T-P interdependence from decadal to multi-centennial timescale. Therefore, we employ paleoclimate proxy data, instrumental datasets and model simulations to obtain a comprehensive view of the relationship between temperature and precipitation changes across Asia.

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75 2. Data

76 2.1. Paleoclimate Data

We identified eleven suitable Holocene paleoclimate proxy reconstructions
for temperature and/or precipitation in the region 60-150°E and 5-50°N after
a quality screening of available data. The datasets cover multiple proxies,
reconstruction techniques and resolutions in the area depicted in Fig. 1.

We only included proxy records which were interpreted as temperature 81 and/or precipitation sensitive by the original authors. Locations, archive and 82 proxy type, seasonal coverage and reconstruction methods of the datasets are 83 given in Table 1. In addition, Table 1 also gives the temporal resolution and 84 the temporal span over which the records were evaluated. The datasets had 85 to cover more than 400 years of the Holocene, between 10 000BP and present 86 day, and had to have sufficient overlap with at least one other complemen-87 tary record. Note that we did not consider individual speleothem $\delta^{18}O$ time 88 series for our analyses, as the attribution to precipitation or temperature 80 may be ambiguous on long timescales [5]. Preliminary analyses indicated 90

that individual cave speleothem time series correlated more strongly with the temperature reconstructions in the set of reconstructions, than with the rainfall reconstructions (not shown). Some speleothem oxygen isotope time series were included in the precipitation reconstruction of Tan et al. [38].

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Reconstruction methods may strongly influence the character and trends 96 of quantitative paleoclimate recostructions, especially when multiple climate 97 variables are derived based on the same proxy data [17, 41]. One tempera-98 ture dataset had to be excluded, because it was by construction negatively 99 correlated to the simultaneous precipitation reconstruction [Number 3 in Ta-100 ble 1, 48]. These climate variables were based on tree-ring and historical 101 drought/flood observation data, and then processed by principle component 102 analyses. Significant axes were combined positively for precipitation, and 103 negatively for temperature. As the temperature reconstruction showed con-104 siderably lower skill than that for precipitation [48] we only retained the 105 summer precipitation time series in the database. 106

Two proxy datasets were considered regionally consistent and comparable, 107 if they both stem from the South-West-Summer-Monsoon (SWSM) domain, 108 west of 100°E, or the East Asian Summer Monsoon (EASM) domain (Fig. 1). 109 A comparison between SWM domain records and EASM records would not 110 be appropriate, as the monsoon systems may act independently and asyn-111 chronously. An independent verification of all proxy reconstructions with 112 meteorological observations is, unfortunately, not possible, as many recon-113 structions do not cover the instrumental era at a sufficient resolution - or at 114 all. 115



Figure 2: Paleoclimate proxy archive composition for each timescale.

117 2.2. Model data

We analyze the climate model simulations from the Coupled Model In-118 tercomparison Project phase 5 (CMIP5) of the last millenium (past1000, 119 850–1850AD) forced with reconstructed solar, volcanic, GHG and aerosol 120 forcing, and partly land use changes [40]. These nine millenium simulations, 121 for which complete surface temperature and precipitation output was avail-122 able, allow us to investigate the modeled relationship $r_{(t,p)}$ in response to 123 largely natural forcing from annual to multidecadal timescales. If multiple 124 ensemble members were available, only the first ensemble member was ana-125 lyzed. To extend our analysis to centennial timescales we employ an orbital 126

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Table 1: Details for paleoclimate reconstructions used in this study. Locations marked with asterisks (*) are based on data from multiple locations, and coordinates in the midst of the representative region are assumed. Sources are given in Supplementary Table 1.

No.	Name	Lat./Lon	Archive	Proxy/ Pa-	Reconstruction	Months	Resolution (Span) [ka BP]	Reference
		$[^{\circ}N/E]$		rameter	method	(temp./p	recip.)	
1	Sheppard-P	35/96	tree	ring width	Temporal re-	1 - 12	.001 (04-2.5)	[35]
				chronol-	gression on			
				ogy/ precip	$\mathrm{met.data}$			
2	Tan 2011 Pr	29/104	$\operatorname{multiple}$	stalagmite	Comparison to	1 - 12	.01 (05-1.9)	[38]
				$\delta^{18} O$ &	model simula-			
				docu-	tion, no cali-			
				ments/	bration			
				precip				
3	NCPrecipIndex	$37/112^{*}$	$\operatorname{multiple}$	tree & doc-	Temporal cal-	1 - 12	.001 (055)	[48]
				uments/	ibration of			
				precip	PCA compo-			
					nents against			
					met.data			
4	Karakoram	36/75	tree	$\delta^{18} {\rm O}/{\rm rainfall}$	Temporal re-	10 - 2	.001 (05-1.0)	[43]
					gression on			
					met.data			
5	Yakumo	42/140	lake	pollen/	Spatial transfer	6,7,8/8	.15 (.04-5.5)	[20]
				pre-	function (Mod-			
				cip&temp	ern Analogs)			
6	Loess	34/110	loess	phytolyths/	Spatial trans-	1 - 12	.25 (2.8-10)	[26]
				pre-	fer function			
				cip&temp	(Weighted Av-			
					eraging Partial			
					Least Squares)			
7	Daihai	41/113	lake	pollen/	Spatial trans-	1 - 12/8	.38 (.01-10)	[47]
				pre-	fer function			
				cip&temp	(Weighted Av-			
					eraging Partial			
					Least Squares)			
8	Sihailongwan	42/126	lake	pollen/	Spatial trans-	1 - 12/8	.06 (.15-10)	[37]
				pre-	fer function			
				cip&temp	(Weighted Av-			
					eraging Partial			
					Least Squares)			
9	PagesAsia2kT	$31/90^{*}$	tree	multiple	Ensemble	1 - 12	.001 (04-1.2)	[30]
					point-by-point			
					regression onto			
					gridded station			
					data			
10	Shihua2003	39/116	stal	layer	Temporal re-	5 - 8	.001 (04-2.6)	[39]
				thickn./temp	gression on			
					met.data			
11	ChinaTempGe	35/110	multiple	multiple/temp	Temporal re-	1 - 12	.01 (025-1.8)	[11]
					gression on			
					met. data			

only forced 6000-year ECHAM5-MPIOM simulation [by 9, denoted "orbital"
in the following]. To cover shorter timescales, and thus to provide a link to
the instrumental record, we employ the 47 historical (hist, 1850-2000 AD)
CMIP5 simulations including natural as well as anthropogenic forcing [40].
A list of the model simulations is provided in Supplementary Table 1.

133 2.3. Instrumental data

Monthly observations from 78 stations were obtained from the Global 134 Historical Climate Network v2 database [GHCN, 31] and averaged to ob-135 tain seasonal values at annual resolution. A year's seasonal average was 136 retained, if joint temperature and precipitation observations were available 137 for all months of the season at the station, and a station time series was 138 considered if at least 50 years of such joint observations could be obtained. 139 Station locations are indicated in Fig. 1, and the number of years each station 140 covers for summer/annual averages is given in Supplementary Table 2. To 141 compare observations to the CMIP5 models, we resampled the 47 CMIP5 his-142 torical model simulations by bilinear interpolation at station/proxy locations 143 and censored the model data to contain the same years as the instrumental 144 records. As a sensitivity test, we additionally analyzed gridded instrumental 145 datasets from the Climate Research Unit (CRU) for precipitation [Hulme 146 Global Land Precipitation Data, ref. 13 and temperature [CRUTEM v.4.2, 147 ref. 16. The gridded datasets were analyzed in the same way, and for the 148 same time periods, as the model data. We focus our discussion on the GHCN 149 instrumental station data, as any processing steps required for deriving the 150 gridded dataset such as infilling of missing data and interpolation influences 151

the precipitation to temperature relationship in unknown ways that are dif-ficult to quantify.

154 3. Methods

155 3.1. Paleoclimate data analysis

To investigate the timescale-dependency of the T-P relationship we low-156 pass filter the data using different cutoff frequencies prior to the correla-157 tion estimation. Orbital-to-millennial scale variability in the Holocene pa-158 leoclimate data was removed by subtraction of a millennial-scale nonlinear 159 trend from a Gaussian kernel smoother with an effective cutoff frequency 160 (halved magnitude of frequency response) $f_{low} = 1/1000$ years. A second 161 Gaussian smoother was passed over the timeseries, with a width given by 162 $f_{hi} = (1/30, 1/50, 1/100...)$ years. This results in a bandpassed time series 163 from which a timescale-dependent Pearson correlation can be estimated ro-164 bustly against irregular sampling of the time series [32]. Significance testing 165 is based on 2000 Monte Carlo simulations using AR1 surrogates with the 166 lag-1 autocorrelation estimated from the f_{low} -detrended proxy time series, 167 and with the original temporal sampling [32]. 168

In a first screening step the original time series had to overlap with more than 50 samples. Then, a pair of records was incorporated for a given timescale of α years if their overlap L was more than 3α , and the mean inter-sample distance was smaller than than 0.5α (Figure 2 and Supplementary Fig. 4). Consider for example two records overlapping over 500 years, at a sample spacing of 10 years for both time series. At a timescale of $\alpha = 100$ years, a cross-correlation would be computed, since the overlap is $5\alpha = 500$ years and the average sample spacing is smaller than $0.5\alpha = 50$ years.

177 3.2. Instrumental data analysis

Both instrumental and model timeseries were linearly detrended and analyzed using timescale dependent Gaussian kernel correlation [32] as for the proxy data. The correlation map fields shown in Fig. 3b were regridded to T63 resolution using bilinear interpolation to allow a comparison of the correlation fields.

183 4. Results

4.1. Multidecadal to centennial-scale temperature to precipitation relation ship

On multi-decadal timescales, the CMIP5 past1000 simulations show a 186 negative summer (JJA) temperature-precipitation correlation (Fig. 3b) over 187 most of Asia's landmass. In particular above peninsular India and Central 188 China, the majority of the individual models agree on a negative correlation 189 sign. The most negative multi-model mean correlations are found above Cen-190 tral India (-0.5) and in Central China (-0.3), the most positive associations 191 above the Pacific, off the coast of Japan (0.6), and in the South China Sea 192 (0.3). The predominant negative relationship over most of the land persists 193 across all testable timescales in summer, winter as well as on the annual 194 mean (Supplementary Figs. 1-3). Above the Tibetan Plateau and the sur-195 rounding ocean basins, the correlation is largely positive, but the model T-P 196 relationship is not as consistent in between the models as above the main-197 land, indicated by fewer plus-signs on Fig. 3b. This may be related to the 198

fact that the representation of the orography of the Himalayan range changes with model resolutions and influences the simulated precipitation [6]. The estimates for the orbital simulation agree well with the multi-model-mean of the past1000 simulation at 30 and 50 year timescale and show that the negative model correlation also persists on decadal to centennial timescales, as shown in Fig. 4.

In contrast, analyzing the proxy data reveals a striking discrepancy be-205 tween the simulated and reconstructed temperature precipitation relation-206 ship: Proxy-based T-P correlations are overall positive, and become stronger 207 at decadal to centennial timescales (Fig. 3). On average, the proxies suggest 208 a positive relationship with $r_{(t,p)} \approx +0.3$ at a 30-year timescale, while the 209 multi-model mean summer (JJAS) relationship at the proxy locations shows 210 an T-P anticorrelation of -0.3. Most (70%) of the individual proxy based 211 correlations are significant at the 90 or 99%-level on centennial timescale, ac-212 counting for autocorrelation as described in Sect. 3.1. The proxy-based T-P 213 correlations are outside the intermodel spread (compare, e.g., Supplemen-214 tary Figs. 1 and 2. for the individual model results on all timescales). Thus, 215 the proxy data suggest that warm centuries tended to be wetter centuries. 216 whereas the simulations suggest that warm centuries corresponded to drier 217 centuries. 218

219 4.2. Annual to decadal scale temperature to precipitation relationship

To test whether a similar difference between models and observations is detectable in the instrumental period, we compare the T-P correlation from GHCN instrumental data and the 47 CMIP5 historical simulations.

²²³ The observed decadal-scale mean correlation, averaged across all stations, be-



Figure 3: Reconstructed and model-based correlation between temperature and precipitation. (a) Timescale dependence of proxy-based correlations. Asterisks indicate significance at the 99%-level, plus-signs at the 90%-level. Green crosses indicate insufficient data. Numbers refer to entries in Table 1, where proxy details are provided. (b) CMIP5 past1000 multimodel mean correlation map between summer (JJAS) temperature and precipitation for the 30-year timescale. Minus (plus) signs indicate, that at least 8 out of the 9 models agree on a negative (positive) sign of the correlation. The color of the circles gives local proxy correlations on the same timescale, dashed lines the correlation between nearby sites.



Figure 4: Timescale-dependent correlation in observations, models and proxy data averaged across station and proxy locations. Symbols and error bars denote the mean correlation between surface temperature and precipitation for the observations and the models and its standard error. The stochastic model (broken lines) fits both proxy and GCM data, with lower temporal persistence (β) and weaker positive coupling at long timescales (b) for the GCMs.

tween summer temperature and rainfall amount is weakly negative ($r_{(t,p)} = -0.13$). For the same locations and years, the simulated decadal-scale multimodel mean correlation at the stations shows a clear negative correlation ($r_{(t,p)} = -0.4$).

The simulated correlation patterns show some model dependence (Fig.3). 228 Subsequently, the model-observation difference in correlation is also model-229 dependent, but most models (37 of 46) reveal a stronger negative average 230 correlation than the observations (Fig. 5). Interestingly, the offset between 231 observed and simulated $r_{(t,p)}$ on decadal timescales is similar to the dis-232 crepancy observed between proxy reconstruction-based and past1000 $r_{(t,p)}$ 233 on multi-decadal timescales (Fig. 4). The direction and magnitude of the 234 model-observation mismatch is therefore consistent across instrumental and 235 proxy data and persists on all analyzed timescales. It is interesting to note 236 that analyzing the CRU gridded datasets instead of the raw station data re-237 sults in a much smaller difference between the observed (gridded) and model 238 based T-P correlations (Fig. 4). Analyzing the gridded data at the station 230 locations, we obtain a decadal-scale T-P correlation of $(r_{(t,p)} = -0.32)$. 240

²⁴¹ 5. Discussion

We observe a clear mismatch between the proxy reconstruction and instrumental based estimates of the T-P relationship and the estimates based on the coupled climate models. This discrepancy may be attributable to weaknesses in proxy reconstructions and instrumental data, deficiencies on the model side, or a combination of both. In the following sections we will explore several potential explanations.



Figure 5: Offset between the CMIP5 historical temperature/precipitation correlations and the station-based correlations on decadal timescales. Critical values of the difference were computed based on the t-statistic with 20 degrees of freedom. Most models display stronger negative correlations than those calculated from the instrumental data.

²⁴⁸ 5.1. Potential reasons for the mismatch on the observational side

There are several levels, at which systematic proxy-dependent effects on the observed T-P relationship could occur.

Firstly, there is a possibility that positive correlations were *induced by* 251 construction during the reconstruction of climate variables, particularly in 252 studies where multiple variables are reconstructed based on the same mul-253 tivariate dataset. This challenge is particularly important for multivariate 254 terrestrial climate archives [17, 41], but also exists for tree-ring-based re-255 constructions [10]. We have excluded one dataset where the correlation was 256 set by construction, as discussed in Section 2.1. If stalagmite δ^{18} O time se-257 ries [e.g. 49] were included as rainfall proxies, the overall correlation would 258 become even more positive (results not shown). We find significant positive 259 correlations (p < 0.01) where temperature and precipitation variables were re-260 constructed from the same dataset and the same methods (Fig. 3). We do, 261 however, also find significant positive correlations (p < 0.05) between rainfall 262 and temperature reconstructions when the proxy, its archive source, and the 263 reconstruction methods differ. 264

Secondly, most proxy reconstructions are subject to considerable uncer-265 tainty concerning their recording season [24] and their recorded climate vari-266 able [17, 41]. Uncertainties with regard to the recording season in proxy 267 reconstructions [24] are not a viable explanation for the model-data discrep-268 ancies, as the model correlation fields remain largely unchanged if annual 269 or boreal winter season mean temperatures and precipitation are considered 270 (Supplementary Figs. 1–3), and the T-P correlations in the GHCN station 271 data are also not sensitive on the analyzed season $(\Delta r_{t,p} = r_{(t,p)}^{\text{JJAS}} - r_{(t,p)}^{\text{ann}})$ is 272

 0.03 ± 0.04). Also, precipitation reconstructions which are directly based on biological archives (i.e. living organisms) may be drought-sensitive, and could reflect variations in soil moisture rather than precipitation. If local temperature and soil moisture in the models were positively correlated, this could close the gap between observations and models. This is, however, not the case as the correlation between temperature and soil moisture is even more negative than the one between temperature and precipitation (Fig. 6).

Thirdly, independent observational noise on the precipitation and tem-280 perature datasets would bias any correlation towards zero and thus lead to 281 underestimation of the reconstructed correlation strength. For the instru-282 mental data this effect should be strongest on interannual timescales as here 283 the relative noise contribution is expected to be highest [18]. This might ex-284 plain why the GHCN station data shows a weaker negative correlation than 285 the gridded CRU data. As a gridcell often averages across multiple stations, 286 this might reduce the observational noise compared to the GHCN station 287 data. On the other hand, such an effect should be reduced on the decadal 288 timescales but we observe a similar model-GHCN observation-gridded CRU 280 observation offset on interannual and decadal timescales. Accounting for 290 noise in the proxy data would even increase the model-data mismatch, as the 291 true underlying positive T-P correlation would then expected to be higher. 292

Another potential reason for the mismatch could be that model data are given as regional (grid-box) averages, while station/proxy data reflect the local climate. Indeed, we find that the correlation in the gridded CRU dataset falls closer to the model data than the GHCN station dataset- shown in Fig. 4. Given that proxy records are often interpreted as reflecting the local climate conditions, we find that it is most appropriate to compare them to station data, instead of grid-box averages. However, we note that the spatial footprint of the specific proxy types, and the dependency of climate variability on the spatial scale, are important open questions for model-data comparisons, which will require further investigation.

Finally, paleoclimate proxy data often contain *temporal uncertainty*. Age uncertainty in one or both of the proxy reconstructions will bias the correlation towards zero [32], contrary to what would be needed to reconcile proxies and observations.

We note that the timescale-dependent change in $r_{(t,p)}$ (Fig. 4) might be 307 affected by the switch of the data type from instrumental data, used up 308 to decadal timescales, to proxy data for longer timescales and the associated 309 changes in the spatial coverage. As we sample the model at the observational 310 sites, the change in spatial coverage does not influence the model-data com-311 parison. To quantify a potential jump in the correlations due to a change in 312 observation locations we employ gridded datasets to compare the correlation 313 difference between proxy and station locations. The T-P mean correlation 314 difference between proxy locations and station locations is small (-0.08 for)315 the CRU gridded dataset, 0.01 for the historical multi-model mean). We 316 therefore conclude that, although there may be a small change due to the 317 location changes, there is no strong influence on the timescale-dependent 318 relationship in Fig. 4. 319

None of described potential shortcomings on the proxy or instrumental data side can readily explain the offset between climate model and paleoclimate reconstruction-based correlation. While given the large number of possible influences upon the proxy reconstruction, we cannot rule out possible biases in the reconstructed relationship, the consistency between instrumental and different proxy evidence suggests that deficiencies in the climate models *or* in the experiment designs (missing or inadequately represented forcings) are at least partly responsible for the mismatch.



Figure 6: Correlation map for the 30-year timescale between summer (JJAS) temperature and precipitation (a) and temperature and soil moisture (b) for the two models (bcccsm-1 and MPI-ESM-P) for which all variables were available at the time of analysis. Minus (plus) signs indicate, that both models agree on a negative (positive) sign of the correlation. Proxy correlations between reconstructed temperature/precipitation are given in both panels as a reference.

³²⁸ 5.2. Potential reasons for the mismatch on climate model side

Mechanisms and processes relating temperature and precipitation variability to each other involve atmospheric dynamics as well as land and ocean surface processes. The offset we observe may be caused by a weakness of positive feedback to balance negative correlations between temperature and precipitation on short timescales, or an overestimation of the negative feedback strength on short timescales.

Stronger negative terrestrial surface T-P correlations in climate models 335 than in the observations have been previously noted on daily [46], monthly 336 [42] and interannual timescales [3]. Several model-based studies have shown 337 that dry (wet) soil tends to suppress (favor) precipitation generation through 338 evapotranspiration decreases (increases), which change local convection, cir-339 culation and moisture advection [see, e.g. 29, 45, 46, and references therein]. 340 Regions such as peninsular India and South-Central China which show dry 341 and warm biases in CMIP5 models [as observed by e.g. 4, 22, 29] may there-342 fore be directly linked to regions with particularly negative T-P relationships, 343 through a negative impact of dryness on moisture advection. Precipita-344 tion underestimation and temperature overestimation on daily to interannual 345 timescales could therefore be a likely candidate for explaining the negative 346 $r_{(t,p)}$ offset. 347

Positive moisture-advection feedbacks relate the monsoon precipitation 348 to a strengthening of the landward circulation, which in turn supplies more 340 moisture. They may dominate the seasonal heat balance on long timescales, 350 and explain abrupt changes in monsoon rainfall under small changes in ex-351 ternal forcing [21]. In the conceptual model of Levermann et al. [21], higher 352 summer temperatures increase the seasonal land-sea thermal contrast due 353 to the different heat capacity of land and ocean, strengthening the monsoon 354 onset circulation and allowing for more precipitation. Weaknesses in the sim-355 ulation of these seasonal processes may explain the lack of positive feedbacks 356 between temperature and rainfall changes on long timescales. 357

There is also considerable *influence of teleconnections and external forcing* on rainfall across Asia, which may modify the T-P relationship on long

timescales. Changes in the frequency of El-Niño/Southern Oscillation events 360 or the states of the Indian Ocean Dipole or the Pacific Decadal Oscilla-361 tion may lead to different temperature and precipitation changes than at-362 mospheric aerosols, solar insolation or greenhouse gas changes [23]. As we 363 analyze fully forced model simulations to derive the correlation structure up 364 to multidecadal timescales and compare the same years of observations and 365 model data in the instrumental period where strong changes in the exter-366 nal forcing occur, different modulations from internal and external forcing 367 should not affect our comparison. The change of determining $r_{(t,p)}$ in the 368 fully forced past1000 simulations (for timescales faster than centennial) vs. 369 in a single orbital-only simulation (for longer timescales) might influence the 370 timescale dependency of the correlation, and leads to a higher uncertainty for 371 the correlation estimate on the long timescales. However, the similar mean 372 of $r_{(t,p)}$ in the orbital only simulation and the fully forced past1000 simula-373 tions further suggests a minor effect of the natural external forcing. This 374 is consistent with the small influence of external forcing on regional climate 375 variability found in [19]. 376

Furthermore, CMIP5/PMIP3 models have been shown to underestimate 377 temperature variability on multidecadal to millennial scales at the sea surface 378 [18, 19], and in the atmosphere [25]. Most immediate and short-term mech-379 anisms influencing the T-P relationship (cloud cover, soil moisture, evapo-380 ration) are expected to induce negative associations. On long timescales, 381 slow-acting components of the earth system (e.g. glaciation changes, basin-382 wide sea surface temperature modes, Intertropical Convergenze Zone shifts) 383 might result in positive links, and could increase the memory of the system 384

³⁸⁵ [25]. However, inadequate representation of such slow feedbacks may not be
 ³⁸⁶ detectable by benchmarking against short observational data.

387 5.3. Comparison using a stochastic model

According to our analysis of proxy and instrumental data, the covariability of temperature and precipitation is timescale-dependent with a negative correlation on annual timescales. On long timescales from multidecadal to centennial, proxy evidence suggests that the relationship is positive. To describe the timescale-dependent behavior and get insights on potential mechanisms, we therefore derive a simple stochastic process-based model with parameters derived from the observed and simulated relationships.

395 5.3.1. Definition

We derive a set of coupled stochastic processes according to a timescale-396 dependent coupling scheme, as illustrated in Fig. 7. 1000-year-long weather-397 and climate-like noise processes W and C are simulated using pink noise pro-398 cesses with a power-spectral density inversely proportional to the frequency, 399 thus following a $1/f^{\beta}$ -behavior [14, 19]. We assume $\beta_1 = 0$ for the high-400 frequency (or short timescale) component W, equivalent to an uncorrelated 401 white noise time series and consistent with weather [12]. For the long-range 402 climate component C we simulate a power-law process with $\beta_2 \in (0, 1)$. The 403 surrogate processes T and P are obtained as weighted means of W and C 404 with weights a and b and c as 405

$$T = aW + bC + cN_1 \tag{1}$$

$$P = -aW + bC + cN_2, \qquad (2)$$

where the sum of weights equals unity, a + b + c = 1. β_1 is kept fixed, while 406 the parameters $\beta_2 = \beta$, a, b and c are varied in the fitting process to mini-407 mize the Root Mean Square Error (RMSE) between observed and simulated 408 correlation between T and P. N_1 and N_2 represents independent observa-409 tion noise on T and P, and is simulated as Gaussian white noise. Pink noise 410 processes were obtained by generating white noise signals, modifying their 411 Fourier transform to obtain the desired slope and re-transforming them into 412 the time domain. The co-variability between T and P was estimated after 413 the removal of the millennial-timescale component, as for the paleoclimate 414 proxy data. 415

Changing the parameters β , a, b and c effectively changes the relationships 416 of the T and P time series. The stochastic model is timescale-dependent, if 417 $\beta > 0$. For example for $\beta = 0$ the obtained time series represent temporally 418 uncorrelated white noise, $\beta \in (0, 1)$ thus allows for varying autocorrelation. 419 Values of β between 1 and 2, on the other hand, would result in time series 420 mimicking glacial-interglacial scale variations [2]. The ratio of a and b de-421 scribes the relative importance of the long-range positive correlation between 422 T and P, whereas parameter c describes independent observational noise. 423

424 5.3.2. Comparison

The parameters of the model were derived by minimizing the RMSE between the correlation estimates from the stochastic model and the climate model and observation-derived correlation estimates from interannual to centennial timescales (the dashed lines in Fig. 4). The best fit to the instrumental and proxy data is obtained for $\beta = 0.5$, a = 0.5, b = 0.4and c = 0.1, which describes a smooth transition from a negative correla-



Figure 7: Illustration of the coupling scheme and spectral characteristics of generating processes in the stochastic model. The final processes T and P are obtained as a linear combination of W, C and N. The influence from process W results in a negative correlation between T and P, that of C in a positive correlation. On short timescales W has a stronger weight, on long timescales C dominates. Additional, mutually independent, observation noise N is added to both T and P.

tion between temperature and precipitation on annual to decadal timescale, 431 to a positive correlation on multidecadal/centennial scale. By contrast, to 432 mimic the covariance simulated by the climate models, the stochastic model 433 needs a weaker timescale dependence ($\beta = 0.25$), a smaller influence of C 434 (b = 0.3) and a stronger noise component (c = 0.2). The lower β of the his-435 torical/past1000/orbital model fit suggests lower climate variability on longer 436 timescales than in the observations. At the same time, positive associations 437 on long timescales are weaker, as b has to be reduced. 438

The stochastic model results support a potential explanation for the discrepancy between models and observations: Weak long-term climate variability in the models, together with less intrinsic positive relationships between local temperature and precipitation, potentially due to soil moisture biases and poor rainfall simulation, yield overall negative associations.

444 6. Conclusion

We have shown that CMIP5/PMIP3 climate model simulations and pale-445 oclimate proxy data suggest considerably different relationships between tem-446 perature and precipitation in Asia on long timescales: Model results suggest, 447 that warmer centuries should have been dryer – proxy results suggest that 448 they were wetter. While we cannot completely rule out systematic biases in 449 the reconstructed T-P relationship, considering the known proxy uncertain-450 ties such as seasonal attribution of the proxy recording system or noise on the 451 proxy records did not resolve the model-data mismatch. Further, indepen-452 dent results such as spatially consistent dry/cold and wet/warm conditions 453 in monsoonal Asia based on quantitative and semi-quantitative moisture in-454

⁴⁵⁵ dicators through the past millennium [7] support our proxy based results.

The observed timescale-dependent nature of the T-P relationship may 456 explain the apparent lack of clear precipitation trends in the past 50 years of 457 the instrumental record [44]: The shortness of the instrumental record only 458 allows to derive synoptic to multidecadal relationships. According to our 459 stochastic model bridging the instrumental and proxy data, negative T-P as-460 sociations dominate at up to decadal timescales and a significant positive re-461 lationship should emerge on considerably longer timescales than 30-50 years. 462 Thus, paleoclimate proxy data may reveal different aspects of the climate 463 system than those emerging from the analyses of the short, high-resolution 464 observational record only. 465

The positive relationship between temperature and precipitation on long timescales in the past may not be directly translated to a warmer and wetter future for Asia as the monsoon response to natural forcing and internal variability in the past may have been different to the response to future increased greenhouse gas emissions [e.g. 23]. However, our results call for a reconciliation of model-data mismatch in the precipitation-temperature relationship which needs attention from both the data and the modeling side.

On the climate model side, the mismatch may be due to intrinsic model aspects, the underestimation of the magnitude of natural forcing, or inadequate sensitivity to forcing [18]. Model sensitivity experiments where parameterizations, forcings or the coupling between components are varied [as e.g. 3, 4, 23] are helpful in this respect, in particular if experiments have been conducted for ensembles, or several models. On centennial timescales, the dynamic adjustment of currently fixed boundary conditions (such as ice sheets and mountain glaciers) may lead to stronger regional variability and
may resolve part of the model-data mismatch we currently observe.

On the proxy side, the uncertainties of reconstructions have to be further 482 explored. To this end, there has been increasing focus on the reconstructabil-483 ity of single or multiple climate variables, in particular from multivariate pale-484 oecological data [17, 41]. Also, spectral biases in proxy archives [as shown for 485 tree-ring data in 10] warrant systematic investigation. Finally, with improved 486 understanding of the processes influencing paleoclimate archives, proxy sys-487 tem models [8] might be developed, which may allow a better comparison of 488 proxy and model (co)variability, and ultimately a resolution of the proxy-data 489 mismatch in the T-P relationship. 490

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Figure (a) Reconstruction-based correlations

(3:11) NC:Ge-(3:10) NC:Shihua-40°N * (8:8) SHL:SHL-* * (7:7) Daihai:Daihai-* * * 30°N (6:6) Loess:Loess-* C (5:5) YAK:YAK-(1:9) Shep:Pages2K-* * * +20°N (4:9) Kara:Pages2K-* * + + + (2:9) Tan:Pages2K-+ (2:10) Tan:Shihua-10°N * (2:11) Tan:Ge 60°E 90°E 120°E 150°E 300 5 10 50 200 30 100 Timescale [years] -1.0 -0.5 0.0 0.5 1.0

(b) Model-based correlations (30 yr. timescale)

Figure 4







